

## Assignment 2 - Some Solutions

- Exercise 1.6.14\*:  $\frac{y}{X} \mathbb{1}_{(X>y)} \leq 1$  and converges to 0 a.s. as  $y \rightarrow \infty$ , so the first result follows from the bounded convergence theorem. To prove the second result, we use our first observation to see that if  $0 < y < \epsilon$ ,

$$E(y/X; X > y) \leq P(0 < X < \epsilon) + E(y/X; X \geq \epsilon).$$

On  $\{X \geq \epsilon\}$ ,  $y/X \leq y/\epsilon \leq 1$  and  $y/X \rightarrow 0$  so the bounded convergence theorem implies

$$\limsup_{y \rightarrow 0} E(y/X; X > y) \leq P(0 < X < \epsilon).$$

Since  $\epsilon$  is arbitrary,

$$\lim_{y \rightarrow 0} E(y/X; X > y) = 0.$$

- Exercise 2.1.7\*:  $E[X_n] = \int_0^1 \sin(2\pi nx) dx = -(2\pi n)^{-1} \cos(2\pi nx)|_0^1 = 0$ . Integrating by parts twice,

$$\begin{aligned} E[X_m X_n] &= \int_0^1 \sin(2\pi mx) \sin(2\pi nx) dx \\ &= \frac{m}{n} \int_0^1 \cos(2\pi mx) \cos(2\pi nx) dx = \frac{m^2}{n^2} \int_0^1 \sin(2\pi mx) \sin(2\pi nx) dx, \end{aligned}$$

so if  $m \neq n$ ,  $E[X_m X_n] = 0 = E[X_m]E[X_n]$ .

To see that  $X_m$  and  $X_n$  are not independent, note that  $X_m(x) = 0$  when  $x = k/2m$ ,  $0 \leq k < 2m$ , and on this set,  $X_n(x)$  takes on the values  $V_n = \{y_0, y_1, \dots, y_{2m-1}\}$ . Let  $[a, b] \subset [-1, 1] \setminus V_n$  with  $a < b$ . Continuity of  $\sin$  implies that if  $\epsilon > 0$  is sufficiently small, we have

$$P(X_m \in [0, \epsilon], X_n \in [a, b]) = 0 < P(X_m \in [0, \epsilon])P(X_n \in [a, b]).$$

- Additional Exercise 1\*

1. We need the following

Lemma: If  $X_n \xrightarrow{P} X$ , as  $n \rightarrow \infty$ , then there exists a subsequence  $X_{n_k}$  such that  $X_{n_k} \rightarrow X$ , almost surely, as  $k \rightarrow \infty$ .

*Proof*. (Idea) We just need to construct a sequence  $\{n_k\}$  such that  $P(|X_{n_k} - X| > 2^{-k}) < 2^{-k}$  (it's easy to show that this is possible. Then, by Borel-Cantelli,

$$P(|X_{n_k} - X| > 2^{-k}, \text{ i.o.}) = 0,$$

which implies that  $X_{n_k} \xrightarrow{a.s.} X$ . □

Since  $X_n$  is increasing, for every  $\omega \in \Omega$ , the sequence  $\{X_n(\omega)\}_{n \geq 1}$  has a limit  $X(\omega)$ . By the lemma, there exist a sequence  $\{n_k\}_{k \geq 1}$  such that the set  $C = \{\omega : X_{n_k}(\omega) \text{ converges}\}$  satisfies  $P(C) = 1$ . By uniqueness of limits, for all  $\omega \in C$ ,  $X_n(\omega) \rightarrow X(\omega)$ .

2. For  $\epsilon > 0$ , define  $A_n^\epsilon = \{\omega : |X_n - X| \geq \epsilon\}$ ,  $A^\epsilon = \limsup_{n \rightarrow \infty} A_n^\epsilon$ . Then

$$\{\omega : X_n \not\rightarrow X\} = \bigcup_{m \geq 1} A^{1/m}.$$

So

$$\begin{aligned} X_n \rightarrow X \text{ a.s.} &\iff P(X_n \not\rightarrow X) = 0 \iff P\left(\bigcup_{m \geq 1} A^{1/m}\right) = 0 \\ &\iff P(A^{1/m}) = 0 \forall m \geq 1 \iff P(A^\epsilon) = 0 \forall \epsilon > 0 \\ &\iff \lim_{n \rightarrow \infty} P\left(\bigcup_{k \geq n} A_k^\epsilon\right) = 0 \iff \lim_{n \rightarrow \infty} P\left(\sup_{k \geq n} |X_k - X| \geq \epsilon\right) = 0 \end{aligned}$$

3. (a) Define  $Y_n = X_1^2 \mathbb{1}\{X_1^2 < \epsilon n^2\}$ . Then  $Y_n \rightarrow X_1^2$ , as  $n \rightarrow \infty$  and  $|Y_n| \leq X_1^2$ . Since  $E[X_1^2] < \infty$  (since  $\text{Var}(X_1) < \infty$ ), we can use the dominated convergence theorem to conclude that

$$\int Y_n dP \rightarrow \int X_1^2 dP. \quad (1)$$

Now if we define  $A_n = \{\omega : |X_1(\omega)| \geq \epsilon \sqrt{n}\}$ , we get

$$E[X_1^2] = \int_{A_n} X_1^2 dP + \int_{A_n^c} X_1^2 dP \geq \epsilon_n^2 P(A_n) + \int Y_n dP.$$

Therefore,  $\epsilon^2 n P(A_n) \leq \int X_1^2 dP - \int Y_n dP$ , which with the help of (1) implies that  $n P(|X_1| \geq \epsilon \sqrt{n}) \rightarrow 0$ , as  $n \rightarrow \infty$ .

- (b) This follows almost immediately from part (a):

$$\begin{aligned} P\left(\max_{1 \leq i \leq n} X_i > \epsilon \sqrt{n}\right) &= P\left(\bigcup_{1 \leq i \leq n} \{X_i > \epsilon \sqrt{n}\}\right) \leq \sum_{i=1}^n P(X_i > \epsilon \sqrt{n}) \\ &= n P(X_1 > \epsilon \sqrt{n}) \leq n P(|X_1| > \epsilon \sqrt{n}) \xrightarrow{n \rightarrow \infty} 0. \end{aligned}$$

4. Let  $\phi(z) = |z|/(1 + |z|)$ .

- (a) Since  $\phi(z) > 0$  for  $z \neq 0$ ,  $E[\phi(|X - Y|)] = 0$  implies  $\phi(|X - Y|) = 0$  a.s. and hence  $X = Y$  a.s.
- (b) It is obvious that  $d(X, Y) = d(Y, X)$ .
- (c) The triangle inequality follows by noting that Exercise 4.10 in Durrett implies that  $\phi(|X - Y|) + \phi(|Y - Z|) \geq \phi(|X - Z|)$  and then taking expected values.

Note that if  $X_n \rightarrow X$  in probability, then since  $\phi \leq 1$ , the dominated convergence theorem implies  $d(X_n, X) = E[\phi(|X_n - X|)] \rightarrow 0$ . To prove the converse let  $\epsilon > 0$  and note that Chebyshev's inequality implies  $P(|X_n - X| > \epsilon) \leq d(X_n, X)/\phi(\epsilon) \rightarrow 0$ .